EBPI Epidemiology, Biostatistics and Prevention Institute

# **Transformation Forests**

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### **Machine Learning**

Machine Learning methods give computers the ability to learn without being explicitly programmed.

(Arthur Samuel, 1959)

Actually: Fit statistical models to data by clever optimisation of appropriate target functions

### **Machine Learning**



Source: https://xkcd.com/1838/

### **Statistical Learning**

An oxymoron, like "Statistical Science"

Either you learn, or you estimate

#### **Statistical Modelling**

Too dull a term to attract any grant money

However: Explicitly acknowledges the underlying probabilistic theory

#### **Statistical Models**

What is a statistical model?

$$Y \sim \mathbb{P}_Y$$

What is a regression model?

$$Y \mid \mathbf{X} = \mathbf{x} \sim \mathbb{P}_{Y \mid \mathbf{X} = \mathbf{x}}$$

#### **Random Forest**

What is a random forest (in general, not only B&C)?

Classical:

$$\mathbb{E}(Y \mid \mathbf{X} = \mathbf{x}) = f(\mathbf{x}), \quad \forall \mathbf{x} \in \mathcal{X}$$

Here:

$$\mathbb{P}(Y \le y \mid \mathbf{X} = \mathbf{x}) = \mathbb{P}_{Y \mid \mathbf{X} = \mathbf{x}}(y) = f(y \mid \mathbf{x}), \quad \forall \mathbf{x} \in \mathcal{X}$$

# Parametric (!) Setup

Unconditional model for response

$$\mathbb{P}_{Y,\Theta} = \{ \mathbb{P}_{Y,\vartheta} \mid \vartheta \in \Theta \}$$

Assumption: Regression model belongs to this family:

$$\mathbb{P}_{Y|\mathbf{X}=\mathbf{x}} = \mathbb{P}_{Y,\vartheta(\mathbf{x})}$$

Task: Estimate  $\vartheta$  function

#### **Likelihood Contributions**

"Learning" data  $(y_i, \mathbf{x}_i), i = 1, \dots, N$  plus family  $\mathbb{P}_{Y,\Theta}$  defines likelihood function

$$\ell_i:\Theta\to\mathbb{R}$$

 $\ell_i(\vartheta(\mathbf{x}_i))$  gives the likelihood for observation i with candidate parameters  $\vartheta(\mathbf{x}_i)$ 

Handle censoring and truncation appropriately here

### **Adaptive Local Likelihood Estimators**

$$\hat{\boldsymbol{\vartheta}}^N(\mathbf{x}) := \argmax_{\boldsymbol{\vartheta} \in \Theta} \sum_{i=1}^N w_i^N(\mathbf{x}) \ell_i(\boldsymbol{\vartheta})$$

Conditioning works via weight functions  $w_i^N(\mathbf{x})$  only

#### **Unconditional Maximum Likelihood**

$$\hat{\boldsymbol{\vartheta}}_{\mathsf{ML}}^{\mathit{N}} := \argmax_{\boldsymbol{\vartheta} \in \Theta} \sum_{i=1}^{\mathit{N}} \ell_i(\boldsymbol{\vartheta})$$

#### **Trees**

$$\mathcal{X} = \bigcup_{b=1,\dots,B}^{\bullet} \mathcal{B}_b$$
 
$$w_{\mathsf{Tree},i}^{N}(\mathbf{x}) := \sum_{b=1}^{B} I(\mathbf{x} \in \mathcal{B}_b \wedge \mathbf{x}_i \in \mathcal{B}_b)$$
 
$$\hat{\vartheta}_{\mathsf{Tree}}^{N}(\mathbf{x}) := \underset{\vartheta \in \Theta}{\mathsf{arg}} \max_{i=1}^{N} w_{\mathsf{Tree},i}^{N}(\mathbf{x}) \ell_i(\vartheta)$$

#### **Forests**

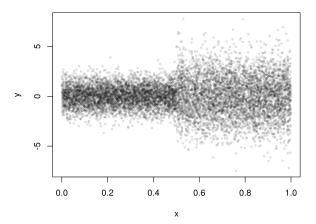
$$\mathcal{X} = igcup_{b=1,\dots,B_t}^{ullet} \mathcal{B}_{tb} ext{ for } t=1,\dots,T ext{ trees}$$
  $w_{\mathsf{Forest},i}^N(\mathbf{x}) := \sum_{t=1}^T \sum_{b=1}^{B_t} I(\mathbf{x} \in \mathcal{B}_{tb} \wedge \mathbf{x}_i \in \mathcal{B}_{tb})$   $\hat{artheta}_{\mathsf{Forest}}^N(\mathbf{x}) := rg \max_{artheta \in \Theta} \sum_{i=1}^N w_{\mathsf{Forest},i}^N(\mathbf{x}) \ell_i(artheta)$ 

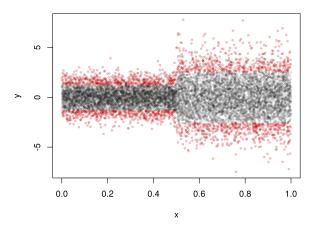
### OK, Done! Really?

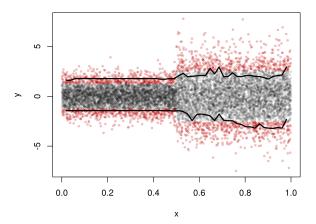
These "nearest neighbor weights" have been used before, first in

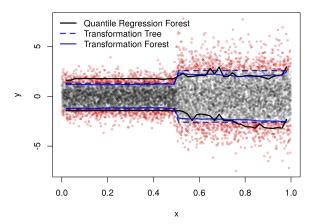
- "bagging survival trees" (2004), in
- "conditional inference forests" (party(kit), since 2005) and in
- "quantile regression forests" (quantregForest, since 2006)
   with standard trees (CART- or CTree-like).

Unfortunately, there is a catch.









#### The Solution

We need splits sensitive to *distributional* and not just *mean* changes.

Generic approach ("Distribution trees and forests"):

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = \mathbb{P}_{Y,\vartheta(\mathbf{x})}(y)$$

Here: Use transformation model

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_Z(\mathbf{a}(y)^\top \vartheta(\mathbf{x}))$$

With

$$\mathbb{P}(Y \le y) = \mathbb{P}(h(Y) \le h(y)) = F_Z(h(y))$$

we can generate *all* distributions  $\mathbb{P}_Y$  from some  $F_Z$  and a corresponding h.

Suitable parameterisations of  $h(y) = \mathbf{a}(y)^{\top} \vartheta$  preserve much of this generality.

As we *always* observe intervals  $(\underline{y}, \overline{y}]$  the exact likelihood is

$$\mathcal{L}(\vartheta|Y \in (\underline{y}, \overline{y}]) := F_Z(\mathbf{a}(\overline{y})^\top \vartheta) - F_Z(\mathbf{a}(\underline{y})^\top \vartheta)$$

- Always defined, always a probability (Lindsey, 1999, JRSS-D)
- Applicable to discrete responses
- Covers all types of random censoring and truncation
- For a precise datum y of some continuous Y, the likelihood can be approximated by the density

$$f_{Y}(y) = f_{Z}(\mathbf{a}(y)^{\top}\vartheta)\mathbf{a}'(y)^{\top}\vartheta$$

Three ways to look at a normal linear model:

1.

$$Y = \alpha + \tilde{\mathbf{x}}^{\top} \boldsymbol{\beta} + \sigma \varepsilon, \quad \varepsilon \sim \mathsf{N}(0, 1)$$

$$\mathbb{E}(Y - \alpha | \mathbf{X} = \mathbf{x}) = \tilde{\mathbf{x}}^{\top} \boldsymbol{\beta}$$

2.

$$\mathbb{P}(Y \le y | \mathbf{X} = \mathbf{x}) = \Phi\left(\frac{y - \alpha - \tilde{\mathbf{x}}^{\top} \boldsymbol{\beta}}{\sigma}\right)$$

3.

$$\mathbb{P}(Y \le y | \mathbf{X} = \mathbf{x}) = \Phi(\tilde{\alpha}_1 + \tilde{\alpha}_2 y - \tilde{\mathbf{x}}^{\top} \tilde{\boldsymbol{\beta}}) \\
\mathbb{E}(\tilde{\alpha}_1 + \tilde{\alpha}_2 Y | \mathbf{X} = \mathbf{x}) = \tilde{\mathbf{x}}^{\top} \tilde{\boldsymbol{\beta}}$$

with 
$$\tilde{\alpha}_1 = -\alpha/\sigma$$
,  $\tilde{\alpha}_2 = 1/\sigma > 0$  and  $\tilde{\boldsymbol{\beta}} = \boldsymbol{\beta}/\sigma$ .

View (3) allows us to see that the normal linear model is of the form

$$\mathbb{P}(Y \le y | \mathbf{X} = \mathbf{x}) = F_Z(h_Y(y) - \tilde{\mathbf{x}}^\top \tilde{\boldsymbol{\beta}}) \\
\mathbb{E}(h_Y(Y) | \mathbf{X} = \mathbf{x}) = \tilde{\mathbf{x}}^\top \tilde{\boldsymbol{\beta}}$$

with  $F_Z$  a cdf of an absolutely continuous rv Z and  $h_Y$  a monotone "baseline transformation function".

With  $F_Z(z) = 1 - \exp(-\exp(z))$  and "unspecified"  $h_Y$  we get the continuous proportional hazards, or Cox, model.

Other choices of  $F_Z$  and  $h_Y$  generate all linear transformation models.

"Linear" transformation models

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_{Z}(\mathbf{a}(y)^{\top} \vartheta - \tilde{\mathbf{x}}^{\top} \beta)$$

"Non-linear" transformation models

$$\mathbb{P}(Y \le y \mid \mathbf{X} = \mathbf{x}) = F_{Z}(\mathbf{a}(y)^{\top} \boldsymbol{\vartheta} - \beta(\mathbf{x}))$$

Conditional transformation models

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_{Z}(\mathbf{a}(y)^{\top} \vartheta(\mathbf{x}))$$

with additive structure of  $\vartheta(x)$  Transformation trees/forests

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_{Z}(\mathbf{a}(y)^{\top} \vartheta(\mathbf{x}))$$

with non-linear structure of  $\vartheta(x)$ 

#### **Parameterisation**

Transformation trees and forests based on parameterisation

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_{Z}(\mathbf{a}_{\mathsf{Bs},d}(y)^{\top} \vartheta(\mathbf{x}))$$

- $-\mathbf{a}_{\mathrm{Bs},d}(y)^{\top}\vartheta(\mathbf{x})$  is a smooth, monotonic Bernstein polynomial of degree d
- -d=1 with  $F_Z=\Phi$  means  $\mathbb{P}_{Y|\mathbf{X}=\mathbf{x}}=\mathcal{N}(\mu(\mathbf{x}),\sigma^2(\mathbf{x}))$
- -d = 5 is surprisingly flexible

# Model-based Recursive Partitioning (MOB)

#### Core idea

- Fit parameters  $\hat{artheta}_{\mathsf{ML}}$  in *unconditional* model  $\mathbb{P}_{Y,artheta}$
- Compute individual gradient contributions ("scores")

$$\mathbf{s}_i = \left. rac{\partial \ell_i(oldsymbol{artheta})}{\partial oldsymbol{artheta}} 
ight|_{oldsymbol{artheta} = \hat{oldsymbol{artheta}}_{\mathsf{ML}}}$$

- Select predictor from  $\mathbf{x}$  with strongest parameter instability as indicated by highest association to  $\mathbf{s}_i$ , i = 1, ..., N
- Find "best" binary split; repeat recursively

Implemented for many models, including (G)LM(M)s, parametric survival,  $\beta$ -regression, spatial lag, Bradley-Terry-Luce, various Item Response Theory models, subgroup analyses, etc.

# **Transformation Trees (TTree)**

- Start with  $\hat{\vartheta}_{\mathsf{ML}}^{N}$
- Search for parameter instabilities in  $\hat{\vartheta}_{\mathsf{ML}}^{N}$  as a function of  $\mathbf{x}$  using (a beefed-up version) of MOB
- Potentially find changes in the mean AND higher moments
- Forests: Aggregate these trees via adaptive local likelihood estimation

# **Transformation Forests (TForest)**

$$\hat{\mathbb{P}}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = \Phi(\mathbf{a}_{\mathsf{Bs},d}(y)^{\top} \hat{\vartheta}^{N}_{\mathsf{Forest}}(\mathbf{x}))$$

makes the forest "parametric" (one model for each x) with

- Forest likelihood
- Prediction intervals
- Likelihood-based variable importance
- Parametric bootstrap
- ...

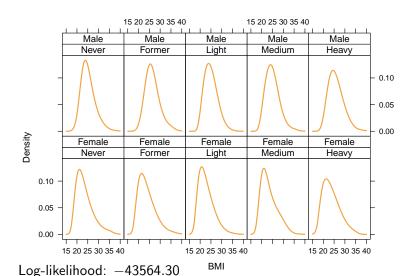
and applicable to censored and truncated data.

# **Swiss Body Mass Index Distributions**

2012 survey (N = 16427) in Switzerland Explain conditional distribution of BMI given

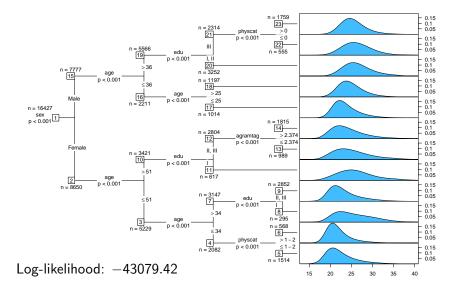
- Sex,
- Smoking status,
- Age,
- Education.
- Physical activity,
- Alcohol intake.
- Fruit and vegetable consumption,
- Region, and
- Nationality.

### **BMI** by Sex and Smoking

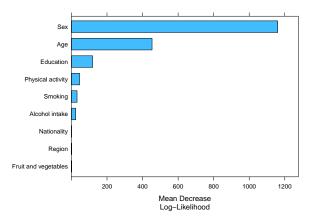


University of Zurich, EBPI Linköpings Universitet, 2018-05-23 Transformation Forests

#### **Transformation Tree**

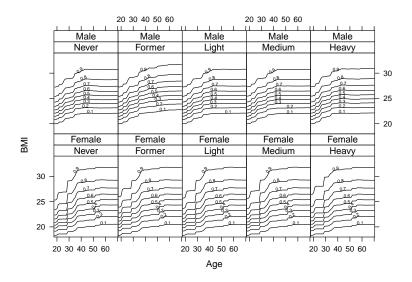


### **Transformation Forest: Variable Importance**



Log-likelihood: -42520.18

#### **Transformation Forest: Partial Deciles**



### **More Complex Models**

For example: Subgroup analysis, stratified / personalised medicine, ...

Conditional transformation model

$$\mathbb{P}(Y \leq y \mid \mathsf{treatment}, \mathbf{X} = \mathbf{x}) = F_Z(\mathbf{a}_{\mathsf{Bs},d}(y)^\top \vartheta(\mathbf{x}) - \beta(\mathbf{x}) I(\mathsf{treated}))$$

- Both the "intercept function"  $\mathbf{a}_{\mathsf{Bs},d}(y)^{\top}\vartheta(\mathbf{x})$  and
- the treatment effect  $\beta(\mathbf{x})$  may depend on  $\mathbf{x}$
- $F_Z() = 1 \exp(-\exp())$  makes  $\beta$  a log-hazard ratio
- Include  $\hat{\beta}$  in search for parameter instabilities

#### **Stratified Medicine**

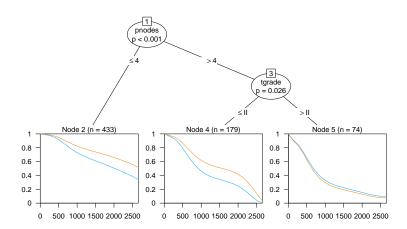
Partition log-hazard ratio  $\beta$  from a fully parametric Cox model

$$\mathbb{P}(T > t \mid \text{treatment}) = \exp(-\exp(\mathbf{a}_{\mathsf{Bs},d}(t)^{\top}\vartheta - \beta I(\text{treated}))$$

for a randomised controlled clinical trial on hormonal treatment of breast-cancer patients

```
> library("tram")
> cmod <- Coxph(ctime ~ horTh, data = GBSG2)
> library("trtf")
> tmod <- trafotree(cmod,
+ formula = ctime ~ horTh | .,
+ data = GBSG2)</pre>
```

#### **Stratified Medicine**



#### **Discussion**

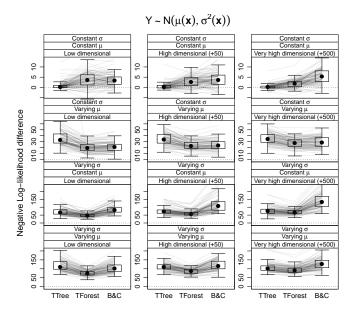
- The "two cultures" of statistical modelling come closer
- With Y = BMI, rain, house prices, survival time etc.

$$\hat{\mathbb{E}}(Y|\mathbf{X}=\mathbf{x}) = \hat{f}(\mathbf{x}) = \mathbf{x}^{\top}\hat{\boldsymbol{\beta}}$$

not interesting (or even harmful)

- $-\mathbb{P}_{Y,\hat{\vartheta}(\mathbf{x})}$  more informative
- Flexibility (non-linear interactions) of B&C random forests preserved
- Simplicity of B&C random forests preserved
- Large sample behaviour?
- High dimensional?

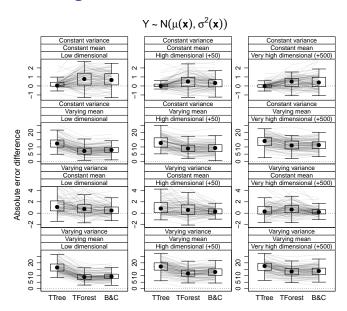
#### Low and High



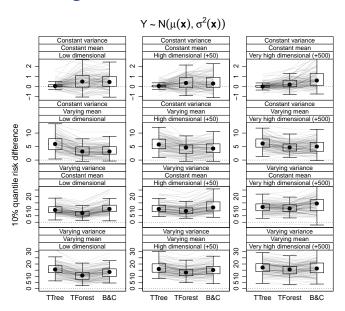
#### Resources

- "Transformation Forests", trtf, https://arxiv.org/abs/1701.02110,
- "Top-Down Transformation Choice" (with BMI example),
   SM, trtf, http://arxiv.org/abs/1706.08269
- "Most Likely Transformations", SJoS, mlt, tram, http://dx.doi.org/10.1111/sjos.12291
- "Conditional Transformation Models", JRSS-B, http://dx.doi.org/10.1111/rssb.12017
- "Model-based Recursive Partitioning", JCGS, partykit http://dx.doi.org/10.1198/106186008X319331,
- "Model-based Recursive Partitioning for Subgroup Analyses", IJB, model4you http://dx.doi.org/10.1515/ijb-2015-0032
- "Model-based Forests", SMMR, model4you, http://dx.doi.org/10.1177/0962280217693034

### Low and High: Median



### Low and High: 10% Quantile



### Low and High: 90% Quantile

