

EBPI Epidemiology, Biostatistics and Prevention Institute

Transformation Forests

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Erasmus University Rotterdam, 2019-03-22

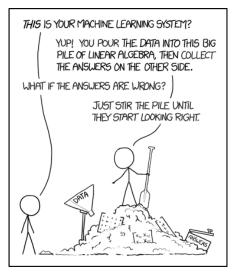
Machine Learning methods give

computers the ability to learn without being explicitly programmed.

(Arthur Samuel, 1959)

Actually: Fit statistical models to data by clever optimisation of appropriate target functions

Machine Learning



Source: https://xkcd.com/1838/

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Statistical Learning

An oxymoron, like "Statistical Science"

Either you learn, or you estimate

Statistical Modelling

Too dull a term to attract any grant money

However: Explicitly acknowledges the underlying probabilistic theory

Statistical Models

What is a statistical model?

 $Y \sim \mathbb{P}_Y$

What is a regression model?

$$Y \mid \mathbf{X} = \mathbf{x} \sim \mathbb{P}_{Y \mid \mathbf{X} = \mathbf{x}}$$

What is a random forest (in general, not only B&C)?

Classical:

$$\mathbb{E}(Y \mid \mathbf{X} = \mathbf{x}) = f(\mathbf{x}), \quad \forall \mathbf{x} \in \mathcal{X}$$

Here:

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = \mathbb{P}_{Y \mid \mathbf{X} = \mathbf{x}}(y) = f(y \mid \mathbf{x}), \quad \forall \mathbf{x} \in \mathcal{X}$$

Parametric (!) Setup

Unconditional model for response

$$\mathbb{P}_{\boldsymbol{Y},\boldsymbol{\Theta}} = \{\mathbb{P}_{\boldsymbol{Y},\boldsymbol{\vartheta}} \mid \boldsymbol{\vartheta} \in \boldsymbol{\Theta}\}$$

Assumption: Regression model belongs to this family:

$$\mathbb{P}_{Y|\mathbf{X}=\mathbf{x}} = \mathbb{P}_{Y,\vartheta(\mathbf{x})}$$

Task: Estimate ϑ function

Likelihood Contributions

"Learning" data $(y_i, \mathbf{x}_i), i = 1, ..., N$ plus family $\mathbb{P}_{Y,\Theta}$ defines likelihood function

 $\ell_i: \Theta \to \mathbb{R}$

 $\ell_i(\vartheta(\mathbf{x}_i))$ gives the likelihood for observation *i* with candidate parameters $\vartheta(\mathbf{x}_i)$

Handle censoring and truncation appropriately here

Adaptive Local Likelihood Estimators

$$\hat{artheta}^{N}(\mathbf{x}) := rg\max_{artheta \in \Theta} \sum_{i=1}^{N} w_{i}^{N}(\mathbf{x}) \ell_{i}(artheta)$$

Conditioning works via weight functions $w_i^N(\mathbf{x})$ only

Unconditional Maximum Likelihood

$$\hat{artheta}_{\mathsf{ML}}^{\mathsf{N}} := rg\max_{artheta \in \Theta} \sum_{i=1}^{\mathsf{N}} \ell_i(artheta)$$

Trees

Forests

$$\begin{split} \mathcal{X} &= \bigcup_{b=1,\dots,B_t}^{\bullet} \mathcal{B}_{tb} \text{ for } t = 1,\dots,T \text{ trees} \\ w_{\text{Forest},i}^{N}(\mathbf{x}) &:= \sum_{t=1}^{T} \sum_{b=1}^{B_t} I(\mathbf{x} \in \mathcal{B}_{tb} \wedge \mathbf{x}_i \in \mathcal{B}_{tb}) \\ \hat{\vartheta}_{\text{Forest}}^{N}(\mathbf{x}) &:= \arg \max_{\vartheta \in \Theta} \sum_{i=1}^{N} w_{\text{Forest},i}^{N}(\mathbf{x}) \ell_i(\vartheta) \end{split}$$

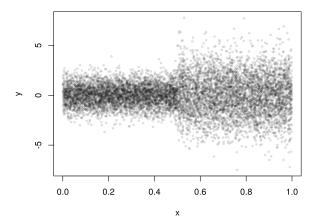
OK, Done! Really?

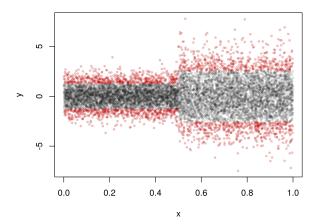
These "nearest neighbor weights" have been used before, first in

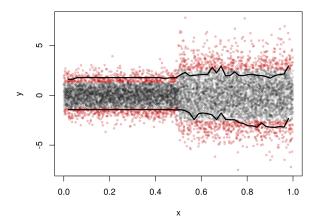
- "bagging survival trees" (2004), in
- "conditional inference forests" (party(kit), since 2005) and in

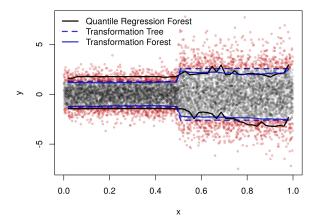
- "quantile regression forests" (**quantregForest**, since 2006) with *standard* trees (CART- or CTree-like).

Unfortunately, there is a catch.









The Solution

We need splits sensitive to *distributional* and not just *mean* changes.

Generic approach ("Distribution trees and forests"):

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = \mathbb{P}_{Y, \vartheta(\mathbf{x})}(y)$$

Here: Use transformation model

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_Z(\mathbf{a}(y)^\top \vartheta(\mathbf{x}))$$

With

$$\mathbb{P}(Y \leq y) = \mathbb{P}(h(Y) \leq h(y)) = F_Z(h(y))$$

we can generate *all* distributions \mathbb{P}_Y from some F_Z and a corresponding *h*.

Suitable parameterisations of $h(y) = \mathbf{a}(y)^{\top} \vartheta$ preserve much of this generality.

As we *always* observe intervals $(\underline{y}, \overline{y}]$ the exact likelihood is

$$\mathcal{L}(\boldsymbol{\vartheta}|\boldsymbol{Y}\in(\underline{y},ar{y}]):= F_{Z}(\mathbf{a}(ar{y})^{ op}\boldsymbol{\vartheta}) - F_{Z}(\mathbf{a}(\underline{y})^{ op}\boldsymbol{\vartheta})$$

- Always defined, always a probability (Lindsey, 1999, JRSS-D)
- Applicable to discrete responses
- Covers all types of random censoring and truncation
- For a precise datum y of some continuous Y, the likelihood can be *approximated* by the density

$$f_{Y}(y) = f_{Z}(\mathbf{a}(y)^{\top}\boldsymbol{\vartheta})\mathbf{a}'(y)^{\top}\boldsymbol{\vartheta}$$

Three ways to look at a normal linear model: 1.

$$Y = \alpha + \tilde{\mathbf{x}}^{\top} \boldsymbol{\beta} + \sigma \varepsilon, \quad \varepsilon \sim \mathsf{N}(0, 1)$$
$$\mathbb{E}(Y - \alpha | \mathbf{X} = \mathbf{x}) = \tilde{\mathbf{x}}^{\top} \boldsymbol{\beta}$$

2.

$$\mathbb{P}(Y \le y | \mathbf{X} = \mathbf{x}) = \Phi\left(\frac{y - \alpha - \mathbf{\tilde{x}}^{\top} \boldsymbol{\beta}}{\sigma}\right)$$

3.

$$\begin{split} \mathbb{P}(\mathbf{Y} \leq y | \mathbf{X} = \mathbf{x}) &= \Phi(\tilde{\alpha}_1 + \tilde{\alpha}_2 y - \tilde{\mathbf{x}}^\top \tilde{\boldsymbol{\beta}}) \\ \mathbb{E}(\tilde{\alpha}_1 + \tilde{\alpha}_2 \mathbf{Y} | \mathbf{X} = \mathbf{x}) &= \tilde{\mathbf{x}}^\top \tilde{\boldsymbol{\beta}} \end{split}$$

with $\tilde{\alpha}_1 = -\alpha/\sigma, \tilde{\alpha}_2 = 1/\sigma > 0$ and $\tilde{\boldsymbol{\beta}} = \boldsymbol{\beta}/\sigma.$

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View (3) allows us to see that the normal linear model is of the form

$$\begin{split} \mathbb{P}(Y \leq y | \mathbf{X} = \mathbf{x}) &= F_Z(h_Y(y) - \tilde{\mathbf{x}}^\top \tilde{\boldsymbol{\beta}}) \\ \mathbb{E}(h_Y(Y) | \mathbf{X} = \mathbf{x}) &= \tilde{\mathbf{x}}^\top \tilde{\boldsymbol{\beta}} \end{split}$$

with F_Z a cdf of an absolutely continuous rv Z and h_Y a monotone "baseline transformation function".

With $F_Z(z) = 1 - \exp(-\exp(z))$ and "unspecified" h_Y we get the continuous proportional hazards, or Cox, model.

Other choices of F_Z and h_Y generate all linear transformation models.

"Linear" transformation models

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_{Z}(\mathbf{a}(y)^{\top} \boldsymbol{\vartheta} - \tilde{\mathbf{x}}^{\top} \boldsymbol{\beta})$$

"Non-linear" transformation models

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_{Z}(\mathbf{a}(y)^{\top} \boldsymbol{\vartheta} - \beta(\mathbf{x}))$$

Conditional transformation models

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_Z(\mathbf{a}(y)^\top \boldsymbol{\vartheta}(\mathbf{x}))$$

with additive structure of $\vartheta(\mathbf{x})$ Transformation trees/forests

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_Z(\mathbf{a}(y)^\top \boldsymbol{\vartheta}(\mathbf{x}))$$

with non-linear structure of $\vartheta(\mathbf{x})$

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Parameterisation

Transformation trees and forests based on parameterisation

$$\mathbb{P}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = F_Z(\mathbf{a}_{\mathsf{Bs},d}(y)^\top \vartheta(\mathbf{x}))$$

- $\mathbf{a}_{\text{Bs},d}(y)^{\top} \vartheta(\mathbf{x})$ is a smooth, monotonic Bernstein polynomial of degree d

$$- d = 1$$
 with $F_Z = \Phi$ means $\mathbb{P}_{Y|\mathbf{X}=\mathbf{x}} = \mathcal{N}(\mu(\mathbf{x}), \sigma^2(\mathbf{x}))$

$$- d = 5$$
 is surprisingly flexible

Model-based Recursive Partitioning (MOB)

Core idea

- Fit parameters $\hat{\vartheta}_{\mathsf{ML}}$ in *unconditional* model $\mathbb{P}_{Y, \vartheta}$
- Compute individual gradient contributions ("scores")

$$\mathbf{s}_i = \left. rac{\partial \ell_i(oldsymbol{artheta})}{\partial oldsymbol{artheta}}
ight|_{oldsymbol{artheta} = \hat{oldsymbol{artheta}}_{\mathsf{ML}}}$$

- Select predictor from **x** with strongest parameter instability as indicated by highest association to $\mathbf{s}_i, i = 1, ..., N$
- Find "best" binary split; repeat recursively

Implemented for many models, including (G)LM(M)s, parametric survival, β -regression, spatial lag, Bradley-Terry-Luce, various Item Response Theory models, subgroup analyses, etc.

Transformation Trees (TTree)

- Start with $\hat{\vartheta}^{N}_{\mathsf{ML}}$
- Search for parameter instabilities in $\hat{\vartheta}_{\rm ML}^N$ as a function of ${\bf x}$ using (a beefed-up version) of MOB
- Potentially find changes in the mean AND higher moments
- Forests: Aggregate these trees via adaptive local likelihood estimation

Transformation Forests (TForest)

$$\hat{\mathbb{P}}(Y \leq y \mid \mathbf{X} = \mathbf{x}) = \Phi(\mathbf{a}_{\mathsf{Bs},d}(y)^{\top} \hat{\vartheta}_{\mathsf{Forest}}^{N}(\mathbf{x}))$$

makes the forest "parametric" (one model for each \mathbf{x}) with

- Forest likelihood
- Prediction intervals
- Likelihood-based variable importance
- Parametric bootstrap

- . . .

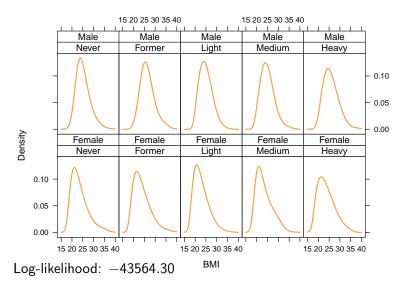
and applicable to censored and truncated data.

Swiss Body Mass Index Distributions

2012 survey (N = 16427) in Switzerland Explain conditional distribution of BMI given

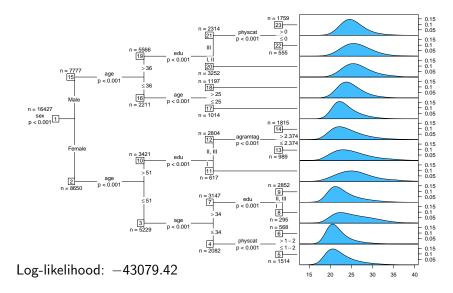
- Sex,
- Smoking status,
- Age,
- Education,
- Physical activity,
- Alcohol intake,
- Fruit and vegetable consumption,
- Region, and
- Nationality.

BMI by Sex and Smoking

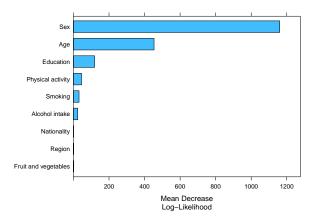


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Transformation Tree



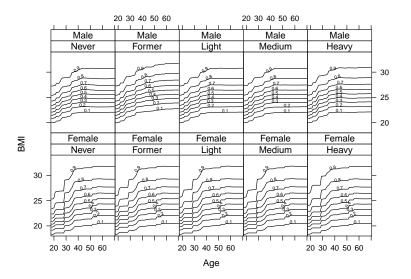
Transformation Forest: Variable Importance



Log-likelihood: -42520.18

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Transformation Forest: Partial Deciles



More Complex Models

For example: Subgroup analysis, stratified / personalised medicine, \ldots

Conditional transformation model

$$\mathbb{P}(Y \leq y \mid \texttt{treatment}, \textbf{X} = \textbf{x}) = \textit{F}_{Z}(\textbf{a}_{\text{Bs},d}(y)^{\top} \boldsymbol{\vartheta}(\textbf{x}) - \beta(\textbf{x})\textit{I}(\texttt{treated}))$$

- Both the "intercept function" $\mathbf{a}_{\mathsf{Bs},d}(y)^{\top} \vartheta(\mathbf{x})$ and
- the treatment effect $\beta(\mathbf{x})$ may depend on \mathbf{x}
- $F_Z() = 1 \exp(-\exp())$ makes β a log-hazard ratio
- Include $\hat{\beta}$ in search for parameter instabilities

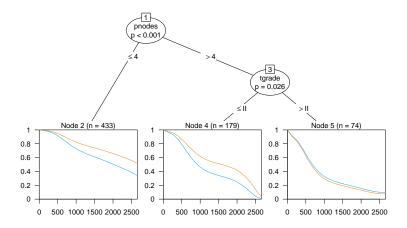
Stratified Medicine

Partition log-hazard ratio β from a fully parametric Cox model

$$\mathbb{P}(\mathsf{\mathcal{T}} > t \mid \texttt{treatment}) = \exp(-\exp(\mathbf{a}_{\mathsf{Bs},d}(t)^\top \boldsymbol{\vartheta} - \beta \mathsf{I}(\texttt{treated}))$$

for a randomised controlled clinical trial on hormonal treatment of breast-cancer patients

Stratified Medicine



Survival Forests

Log-rank splitting implicitly assumes proportional hazards model

$$\mathbb{P}(T > t \mid \mathbf{X} = \mathbf{x}) = \exp(-\exp(h(y) - \beta(\mathbf{x})))$$

 \Rightarrow cforest, ranger, randomForestSRF are insensitive to non-proportional hazards effects.

Switching to transformation forests based on

$$\mathbb{P}(T > t \mid \mathbf{X} = \mathbf{x}) = \exp(-\exp(\mathbf{a}(y)^{\top}\vartheta(\mathbf{x})))$$

relaxes this restriction.

Discussion

- The "two cultures" of statistical modelling come closer
- With Y = BMI, rain, house prices, survival time etc.

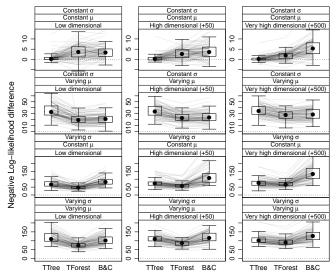
$$\hat{\mathbb{E}}(Y|\mathbf{X} = \mathbf{x}) = \hat{f}(\mathbf{x}) = \mathbf{x}^{\top}\hat{\boldsymbol{\beta}}$$

not interesting (or even harmful)

- $\mathbb{P}_{\mathbf{Y}, \hat{\boldsymbol{\vartheta}}(\mathbf{x})}$ more informative
- Flexibility (non-linear interactions) of B&C random forests preserved
- Simplicity of B&C random forests preserved
- Large sample behaviour?
- High dimensional?

Low and High

$Y \sim \mathsf{N}(\mu(\boldsymbol{x}), \sigma^2(\boldsymbol{x}))$

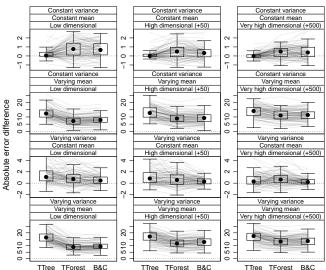


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Resources

- "(Survival) Transformation Forests", trtf, https://arxiv.org/abs/1701.02110, https://arxiv.org/abs/1902.01587
- "Top-Down Transformation Choice" (with BMI example), SM, trtf, http://arxiv.org/abs/1706.08269
- "Most Likely Transformations", SJoS, mlt, tram, http://dx.doi.org/10.1111/sjos.12291
- "Conditional Transformation Models", JRSS-B, http://dx.doi.org/10.1111/rssb.12017
- "Model-based Recursive Partitioning", JCGS, partykit http://dx.doi.org/10.1198/106186008X319331,
- "Model-based Recursive Partitioning for Subgroup Analyses", IJB, model4you http://dx.doi.org/10.1515/ijb-2015-0032
- "Model-based Forests", SMMR, model4you, http://dx.doi.org/10.1177/0962280217693034, AOAS https://arxiv.org/abs/1804.02921

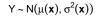
Low and High: Median

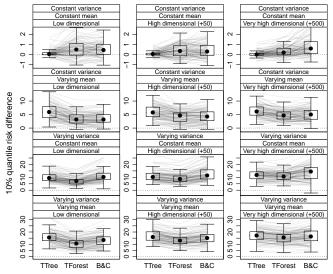


 $Y \sim N(\mu(\mathbf{x}), \sigma^2(\mathbf{x}))$

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Low and High: 10% Quantile

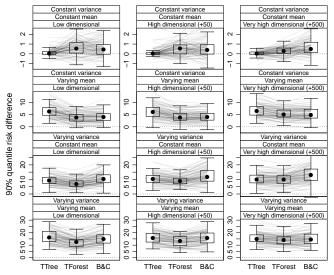




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Low and High: 90% Quantile

 $Y \sim N(\mu(\mathbf{x}), \sigma^2(\mathbf{x}))$



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